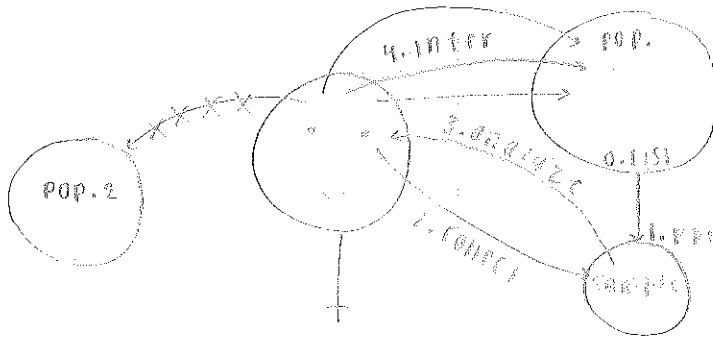


hypothesis testing

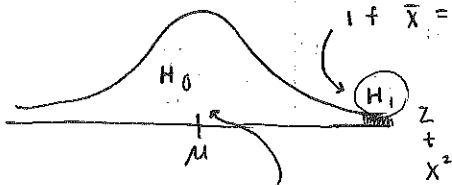
one sample



- **hypothesis testing**: a procedure, based on sample info by which one "accepts" or "rejects" the hypothesis
↳ "fail to reject"
 - **Null hypothesis**: H_0 the hypothesis set up to test against
- "there is no change / difference" ↳ the standard
 - **Alternate hypothesis**: H_1 the hypothesis to be accepted if the the null is rejected ↳ the alternative
- "there is a difference"
- 3 looks of the alternate hypothesis
↳ one-tailed test to the right if $\mu \geq 100$, critical region on right
↳ one-tailed test to the left if $\mu \leq 100$, critical region on left
↳ two-tailed test if $\mu \neq 100$, has critical region on both sides
- * μ will not always be 100, this was just an example *

$$H_0: \mu = 100$$

$$H_1: \mu \neq 100$$



if $\bar{x} = 100.57, z = .123$ null hypothesis is not rejected

- **hypothesis testing errors**
 - **type I error**: rejecting the null hypothesis when it's true
↳ worst error; " α " = level of significance = .1, .05, .01
 - **type II error**: accepting the null when it is false
↳ " β "

increase sample size = increase power
similar to probability

	"fail to reject"	"reject"
H_0 : true	☺	type I $\alpha = .1, .05, .01$
H_0 : false	type II β	power of a test $(1 - \beta)$

uncomfortable but correct

4 ingredients for a statistical test:

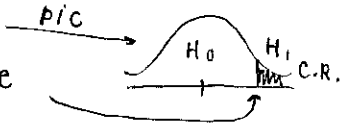
1. H_0 $\mu = .97$

2. H_1 $\mu > .97$

3. critical value

$z_c = 1.645, 1.96...$

4. $\hat{p} \rightarrow z =$



also p-value, but not necessary

print out project checklist

don't need samples / graphs BUT do bivariate data
 ↳ scatterplot...

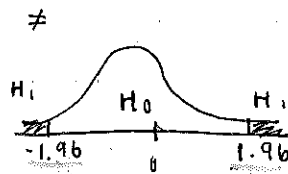
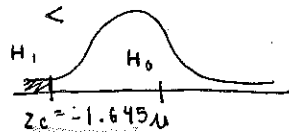
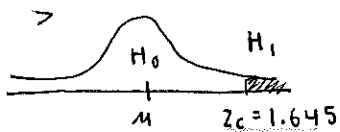
project due end of April

9.2 (11-14, 16) print template

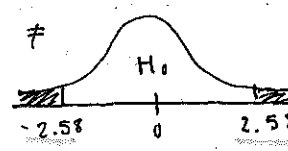
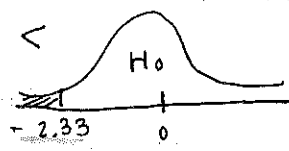
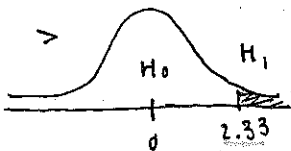
SOH - CAH - TOA

"five is one and one is two"

$\alpha = .05$



$\alpha = .01$ ← usually for risky topics



test μ when σ is known

SRS, get \bar{x}

ex. sun spots $\bar{x} = 47.0$

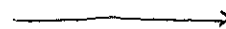
"higher than"

$H_1: >$

prev. studies $\sigma = 35$ $\mu = 41$
 avg. from thousands
 of years

default

$\alpha = .05$ $n = 40$



inference template


name of test

1-sample mean z-test

$H_0: \mu = 41$ sun spot cycles have an avg of 41 per cycle

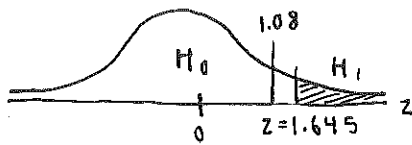
$H_1: \mu > 41$

conditions:

- SRS
- independent
- $n = 40 > 30 \therefore$ CLT invoked \rightarrow 
- σ known

"it's reasonable to assume blah blah blah"


pic.:



$$z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} \quad \bar{x} = 47 \quad z = \frac{47 - 41}{35/\sqrt{40}} = 1.08$$

conclusion:

- r $\textcircled{H_0}$ ~~H_1~~ $\alpha = .05$

-  \exists insufficient statistical evidence suggesting there are higher than 41 sun spots $\therefore \mu \approx 41$

ex. sprinkler system for fires

$\mu = 130^\circ\text{F}$ $n = 81$ $\bar{x} = 131.08^\circ\text{F}$ $\sigma = 1.5^\circ\text{F}$ a contradiction of H_0 ?

$H_0 = 130$

$H_1: \mu > 130$

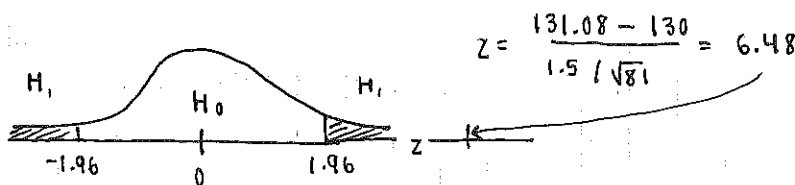
$\alpha = .05$

\neq

1-sample mean test w/ z curve

$H_0: \mu = 130^\circ\text{F}$ Sprinklers are activated at an avg of 130°F


$H_1: \mu \neq 130^\circ\text{F}$




$$z = \frac{131.08 - 130}{1.5/\sqrt{81}} = 6.48$$

conclusion:

- r ~~H_0~~ $\textcircled{H_1}$ $\alpha = .05$

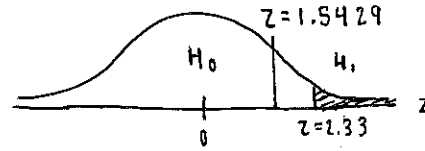
-  \exists statistical evidence suggesting that the sprinklers do activate at a temp. different than 130°F also $> 130^\circ\text{F}$

9.2 #11 $\mu = 16.4$ ft $n = 36$ $\bar{x} = 17.3$ $\sigma = 3.5$ ft $\alpha = .01$
 is it increasing? $H_1: \mu > 16.4$ ft

- \sqrt{s}
- RRS
- indep.
- $n = 36 > 30 \therefore$ CLT invoked
- σ known,  z

$H_0: \mu = 16.4$ ft
 $H_1: \mu > 16.4$ ft

P
 value



$$z = \frac{17.3 - 16.4}{3.5 / \sqrt{36}} = 1.5429$$

$\textcircled{H_0}$ $\alpha = .01$

$\exists x$ insufficient evidence suggesting the storm is increasing above severe rating of 16.4 ft $\therefore \mu \approx 16.4$ ft

p-value shows prob. of getting results if the null is true
 $\mu = 68.7g$ $\mu = 70g$ $p = .18$ $\alpha = .05$

$p > \alpha$ so accept null

$H_0: \mu = 70g$ 18% chance of getting this result (mean of 68.7 or less) from a sample of this size and variation

education, psych, sociology: $\alpha = .20, .10$ bc less serious

if p is small, there is sufficient evidence against null

read stat 9.2 peck 9.2 add p-values

ex. pepperdine cafe

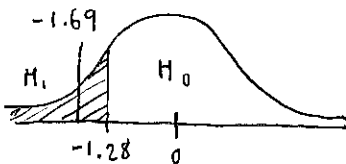
$\mu = 41$ min $n = 40$ $\bar{x} = 38$ min $\sigma = 11.2$ min $\alpha = .10$

$H_0: \mu = 41$ min avg time spent in Pepp. cafe was 41 min at lunch

$H_1: \mu < 41$ min

use z-table $z = -1.28$ shows $\alpha = .10$

or t-table with one-tail



$Z = SNC$

$$z = \frac{38 - 41}{11.2 / \sqrt{40}} = -1.69 \quad p = .0455$$

$\textcircled{H_1}$ $\alpha = .10$

$p < \alpha$, strong evidence against the null

$\exists x$ sufficient statistical evidence that pepp. students spend < 41 min.

testing μ when σ unknown

- get \bar{x} and s from a sample
- 1. H_0, H_1, α
- 2. $t = \frac{\bar{x} - \mu}{s/\sqrt{n}}$ with degrees of freedom $n-1$
- 3. use t table to find p -value

9.2(17-20, 22)

checks:

independent

RRS

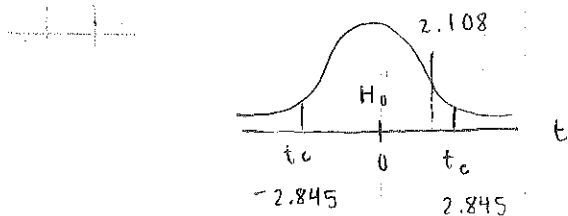


$n < .1N \rightarrow$ "it is reasonable to assume ..."

ex. leukemia drugs

$\bar{x} = 17.1$ $s = 10.0$ $\mu = 12.5$ $\alpha = .01$ $n = 21$

$t = 2.108$



$\bar{x} \rightarrow t = \frac{\bar{x} - \mu}{s/\sqrt{n}}$ (t-score)

$t_c = 2.845$

9.3(9-12, 18-20)

H_0 ~~$\mu = 12.5$~~ $\alpha = .01$

\exists insufficient statistical evidence suggesting that the mean remission time of 6-MP is different from 12.5 weeks $\therefore \mu \approx 12.5$

go to t table looking for $t = 2.108$ $\therefore p$ -value between .05 and .02

$\alpha = .02 < p \text{ value} < .05 \rightarrow .01 < .02 < p \text{ value} < .05 \rightarrow$ fail to reject null
"weak evidence against null"

25. $c =$ confidence level, $\alpha =$ significance level for a two-tailed test

null hypothesis $H_0: \mu = k$

reject H_0 whenever k falls outside of $c = 1 - \alpha$

fail to reject if in the interval $c = 1 - \alpha$

$H_0: \mu = 20, H_1: \mu \neq 20$ $n = 36$ $\bar{x} = 22$ $\sigma = 4$ $\alpha = .01$

a) $c = 1 - \alpha = 1 - .01 = .99$

$ME = z_c \left(\frac{\sigma}{\sqrt{n}} \right) = 2.58 \left(\frac{4}{\sqrt{36}} \right) = 1.72$

$22 - 1.72 < \mu < 22 + 1.72 \rightarrow 20.28 < \mu < 23.72$

$k = 20$, not in interval: reject

b) $Z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} = 3.0000$ $p \text{ value} \approx .0026 < \alpha = .01$, reject null

1 sample proportion z-test

"binomial"

$$np > 5 \text{ and } nq > 5 \rightarrow n(\pi) > 5 \quad n(1-\pi) \geq 10$$

$$\hat{p} = \frac{r}{n} \rightarrow Z = \frac{\hat{p} - \pi}{\sqrt{\pi(1-\pi)/n}}$$

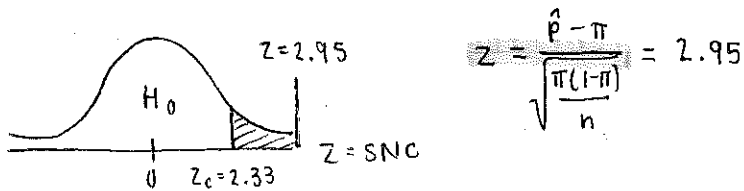
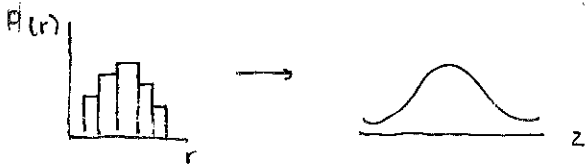
ex. eye surgery

$$\pi = .30 \quad n = 225 \quad r = 88 \quad \hat{p} = 88 / 225 = .3911 \quad \alpha = .01$$

$$H_0: \pi = .30$$

$$H_1: \pi > .30$$

RRS, independent, $n\pi > 5$ $n(1-\pi) > 5$
 $\hookrightarrow 67.5 > 5$ $\hookrightarrow 157.5 > 5 \quad \therefore$ ANALY



$p\text{-value} = .0016 < .01 \rightarrow$ strong evidence against the null

$$H_0 \quad (H_1) \quad \alpha = .01$$

\exists strong statistical evidence suggesting a higher rate of restoration

if $n\pi > 5$ but not > 10 , say "possibly concerned that sample is not large enough"

10.1 (6, 7, 9, 12, 13, 15)

9 review; mileage

$$\mu = 11.1 \text{ thousand } 11,100$$

$$\sigma = 600$$

$$n = 36$$

$$\bar{x} = 10.8 \quad 10,800$$

$$\alpha = .05$$

"different from" \neq

stat \rightarrow calc \rightarrow z-test

$$x \rightarrow z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} = \frac{10.8 - 11.1}{6/\sqrt{36}} = -3$$

$$z_c = 1.96 \text{ and } -1.96$$

p-value $<$ α "SEVN"

$$.0027 < .05$$

large z-value, small p-value

$$\uparrow n \downarrow \sigma \rightarrow \uparrow z_T \text{ or } \downarrow p$$

$$\mu = .35$$

$$n = 81$$

$$r = 39$$

$$\hat{p} = .4815$$

"more than"

$$\alpha = .05$$

$$p\text{-value} < \alpha$$

$$.0066 < .05$$

"SEVN"

$$\mu = 48$$

$$n = 10$$

$$\bar{x} = 44.2$$

$$s = 8.61$$

$$\alpha = .05$$

"less than"

$$t = -1.3952$$

$$p = .0982$$

$$p\text{-value} > \alpha$$

$$.098 > .05$$

"WEVN"

hypothesis testing

two samples

hypothesis testing, dependent groups

paired data samples that help draw conclusions about the difference of 2 groups

before / after, matching even without before / after

• using matched data reduces extraneous factors; reduces variance, increases accuracy

\bar{d} (basically \bar{x}): mean difference between the paired data

S_d : sample standard deviation

$$L \rightarrow \sqrt{\frac{\sum (d - \bar{d})^2}{n-1}}$$

2 dep. mean t-test

$H_0: \mu_d = 0$ "no difference"

$$\bar{d} \rightarrow t = \frac{\bar{d} - \mu_d}{S_d / \sqrt{n}}$$

$H_0: \mu_d = 0$ avg diff. of creativity training - control 20

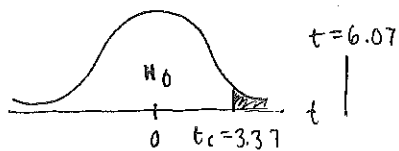
$H_1: \mu_d > 0$

$\alpha = .01$

$n = 6$

$df = 5$

$t_c = 3.365$



$$\bar{d} \rightarrow t = \frac{\bar{d} - \mu_d}{S_d / \sqrt{n}} = \frac{9.83 - 0}{3.97 / \sqrt{6}} = 6.06 \approx 6.07$$

~~H_0~~ (H_1) $\alpha = .01$

$P(V|U) < \alpha$ "SEVN"
 $\leq 0.005 < .01$

checks:

RRS

independent / dependent

↳ all independent ↳ within each thing being tested from one another

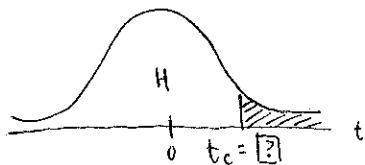
$n < .1N$ "it is reasonable to assume..."

$$L \rightarrow \underbrace{\quad} \rightarrow \underbrace{\quad}_t$$

baby IQ example:

$H_0: \mu_d = 0$ no diff between age 3 and age 8

$H_a: \mu_d > 0$



p value $<$ α "SEVN"
.001 $<$.05

$$t = 3.17$$

confidence interval: $\bar{d} - E < \mu_d < \bar{d} + E$

$$9.91$$

$$2.39 < \mu_d < 17.43$$

$$E = t_c \left(\frac{s_d}{\sqrt{n}} \right)$$

$$E = t_c \frac{22.11}{\sqrt{50}} \quad 98\% \quad t_c \approx 2.412$$

$$E = 7.52$$

population average difference ... 98 times

hypothesis testing σ unknown

2 independent sample means t-test

checks:

2 independent, random, n_1 and $n_2 \geq 30$ \therefore CLT $n_1 < .1N_1$ and $n_2 < .1N_2$
degrees of freedom: σ_1 and σ_2 unknown

- $n_1 - 1$ and $n_2 - 1 = n_1 + n_2 - 2$
- smaller $n - 1$
- satterthwaite's formula

Null / Alternate Hypothesis

$H_0: \mu_1 = \mu_2$ or $\mu_1 - \mu_2 = 0$ identify μ_1 and μ_2

$H_1: \mu_1 < \mu_2$

$\mu_1 > \mu_2$

$\mu_1 \neq \mu_2$

$$t = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}$$

use t-table for this

1. find d.f. and then t_c

confidence interval

$$(\bar{x}_1 - \bar{x}_2) - E < \mu_1 - \mu_2 < (\bar{x}_1 - \bar{x}_2) + E$$

$$E = t_c \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$$

if only negative values, $\mu_1 < \mu_2$

if only positive, $\mu_1 > \mu_2$

if both, no conclusion

$\hookrightarrow \mu_1 \approx \mu_2$

ex. 5

advil: $\bar{x}_1 = 20.1$ min $s_1 = 8.7$ $n_1 = 12$

tylenol: $\bar{x}_2 = 11.2$ min $s_2 = 7.5$ $n_2 = 8 \rightarrow$ d.f. = 7

$H_0: \mu_1 = \mu_2$ the time it takes to enter the bloodstream is the same

$H_1: \mu_1 > \mu_2$

$$t = 2.437 \quad t_c = 1.895$$

$$E = 1.895 \sqrt{\frac{8.7^2}{12} + \frac{7.5^2}{12}} = 6.92$$

$$.010 < p\text{value} < .025 < \alpha = .05 \text{ "SEVN"}$$

$1.98 < \mu_1 - \mu_2 < 15.82$ "if we took..."
 $+ \rightarrow + \Rightarrow \mu_1 > \mu_2$ population difference of means $\mu_1 - \mu_2$

2 independent sample proportion z test

$n_1, r_1, p_1 \quad n_2, r_2, p_2$

for large samples, distribution is approximation of normal

$$\hat{p}_1 - \hat{p}_2 = \frac{r_1}{n_1} - \frac{r_2}{n_2}$$

$H_0: p_1 = p_2$ "no difference"

$H_1: p_1 < p_2$

$p_1 > p_2$

$p_1 \neq p_2$

$$\bar{p} = \frac{r_1 + r_2}{n_1 + n_2} \quad \bar{q} = 1 - \bar{p}$$

checks:

RRS

Indep.

$n_1 < .1N_1, n_2 < .1N_2$

$n_1 \bar{q} > 5$

$n_1 \bar{p} > 5$

$n_2 \bar{q} > 5$

$n_2 \bar{p} > 5$

ANAIJ

$$Z = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\frac{\bar{p}\bar{q}}{n_1} + \frac{\bar{p}\bar{q}}{n_2}}}$$

ex.7 Voting

$n_1 = 625 \quad r_1 = 295$

$p_1 = .472$

$\alpha = .05$

$n_2 = 625 \quad r_2 = 350$

$p_2 = .56$

$H_0: p_1 = p_2$

$H_1: p_1 < p_2$

$Z = -3.11$

α

$p = .0009 < .05 \quad \therefore$ "SEVN"

$$\hat{\sigma} = \sqrt{\frac{\hat{p}_1 \hat{q}_1}{n_1} + \frac{\hat{p}_2 \hat{q}_2}{n_2}}$$

find z_c for confidence interval

$$\text{Error} = z_c \cdot \hat{\sigma}$$

$$(\hat{p}_1 - \hat{p}_2) - E < p_1 - p_2 < (\hat{p}_1 - \hat{p}_2) + E$$

pop. difference of proportions $p_1 - p_2$

10.3 (4, 9, 11, 12, 17, 22)