

7.1 Confidence Intervals

C-Int	Zc
80%	1.28
90%	1.645
95%	1.96
99%	2.58

1) $\epsilon = Z_c \left(\frac{\sigma}{\sqrt{n}} \right)$
maximum error

2) $\bar{x} - \epsilon < \mu < \bar{x} + \epsilon$

(the interval will be larger for higher Percents confidence)

$\uparrow w = \uparrow c$

also $\uparrow n \uparrow c \downarrow w$

Example: 15.6 is μ , σ is 1.8. 95% confidence interval?
 $\epsilon = 1.96 \left(\frac{1.8}{\sqrt{90}} \right)$, $\epsilon = 0.37$
↑
Z for 95%

when pop ↑, the width ↓

$\bar{x} - \epsilon < \mu < \bar{x} + \epsilon$
 $15.6 - 0.37 < \mu < 15.6 + 0.37$
 $15.23 < \mu < 15.97$

99% confidence interval?
 $\epsilon = 2.58 \left(\frac{1.8}{\sqrt{90}} \right)$, $\epsilon = 0.4895$
 $15.6 - 0.49 < \mu < 15.6 + 0.49$
 $15.11 < \mu < 16.09$

★ include on hw + test

conclusion: if we took 100 samples of $n=90$, we expect to catch the population mean (μ) of Julia's 2 mile jogging times 99 occasions

Prerequisites for this to work:

Checks:

- 1) σ known
- 2) RRS
- 3) Independent
- 4) normally distributed or $n > 30$ (CLT)

★ include on hw + test

↑ ↑
 or many samples... 99% occasions

Finding Sample Size

-knowing ϵ , finding n

$n = \left(\frac{Z_c \sigma}{\epsilon} \right)^2$ or plug back in $\epsilon = Z_c \left(\frac{\sigma}{\sqrt{n}} \right)$

$w = 2\epsilon$
 $\epsilon = w/2$

T-Intervals

A24

Tests of Significance / T-table

due to chance or something else?

T-procedures: θ unknown

• Z formulas for significance: $Z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}}$, $X \pm Z_c \left(\frac{\sigma}{\sqrt{n}}\right)$ ^{C for confidence interval}

• if no σ , use S (sd for sample); but w/ this switch we can't use Z , so we use T !

$$-t = \frac{\bar{x} - \mu}{S/\sqrt{n}} \quad , \quad X \pm T \left(\frac{S}{\sqrt{n}}\right)$$

significance confidence

T-distribution

- broader, tails are broader (\uparrow probability of getting results farther from 0, as S varies + more uncertain than σ)

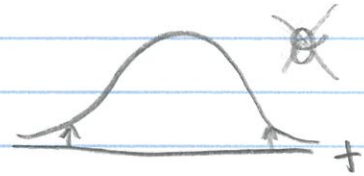
- diff T-distr for every sample size

- as sample size \uparrow , the sd (S) gets closer to σ , big samples more accurate

- \uparrow sample gets closer to normal curve
 \downarrow inside of same sample size

Degrees of Freedom

- sample size minus 1 ($n-1$)
as $df \uparrow$, t distr becomes more normal



WS Gossett

1908, Ireland

Guinness Brewery

chemist

t -distribution

ex: confidence level notation

$$287.7 < \mu < 290.3 \text{ PCB levels}$$

If we took 100 samples of sample

size $n=10$, we expect to capture the pop'n

mean (μ) of PCB levels 95 times

Steps

Go to T-table, down using degrees of freedom and the row is confidence levels (or area under tail.) That gives you T_c (which plug into equation to get confidence interval)

$$E = t_c \left(\frac{S}{\sqrt{n}}\right)$$

Checks

1) RRS

2) Independent

3) normally distributed (approx)

or $n > 30$ or NQP is

linear / pearson's index

7.3

based off binomials: $\frac{r}{n} = \hat{p}$ ← sample percent / proportion

$$\hat{p} = r/n$$

$$\cdot \hat{E} = z_c \left(\sqrt{\frac{\hat{p}\hat{q}}{n}} \right)$$

$$\cdot \hat{p} - \hat{E} < \bar{p} < \hat{p} + \hat{E}$$

- p/π is population % / proportion

- \hat{p} is sample % / proportion

• Checks

1) RRS

2) indep

3) $np_{\text{and } nq} \geq 10$

"ANAD"

$$\frac{7.1}{n = \left(\frac{z_c}{\hat{E}} \right)^2}$$

round up!

estimating n

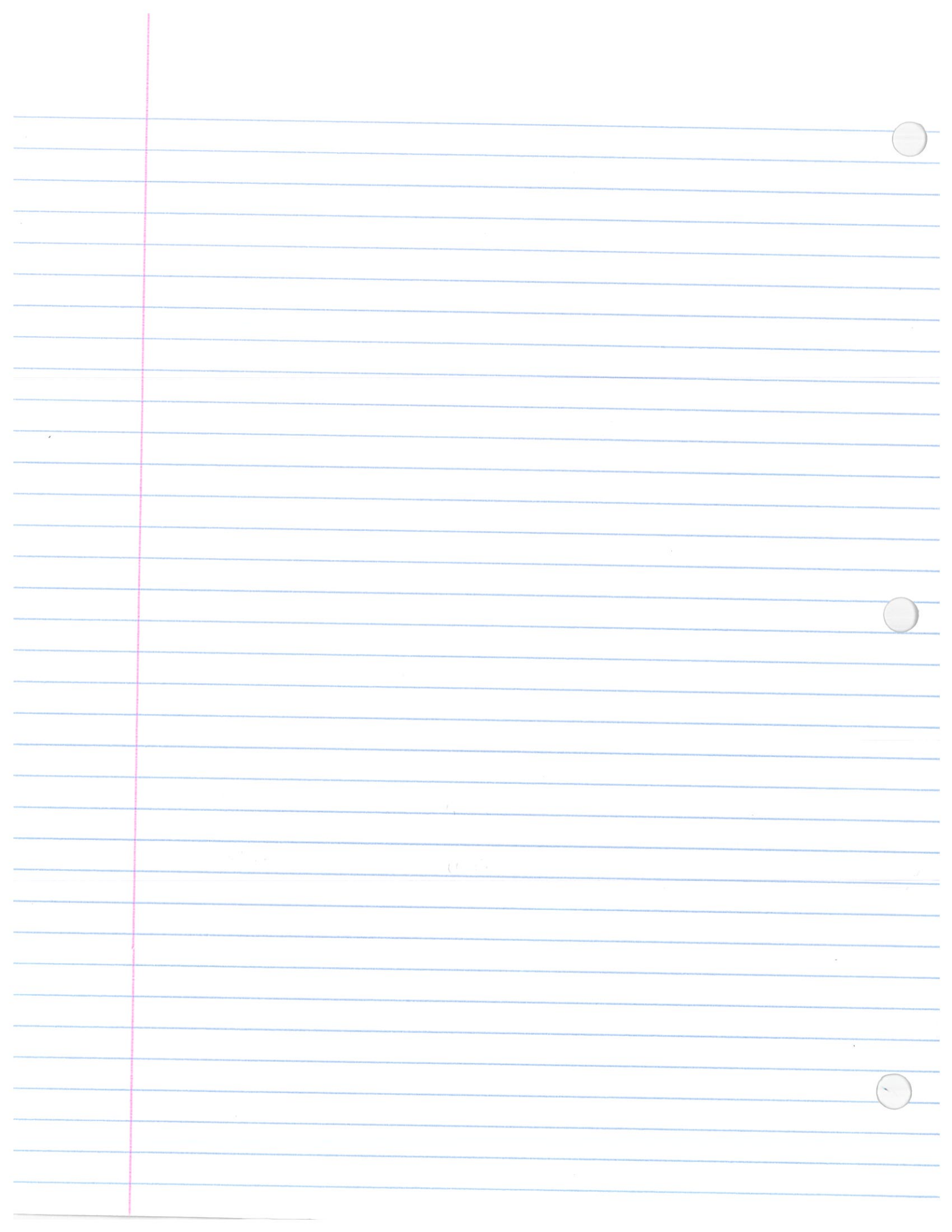
$$\frac{7.3}{n = pq \left(\frac{z_c}{\hat{E}} \right)^2}$$

preliminary estimate
smaller n

Or

$$n = .25 \left(\frac{z_c}{\hat{E}} \right)^2$$

Starting from scratch,
n bigger



7.1 (CI w/ θ)


Z_c	CI%	$E = Z_c \left(\frac{\theta}{\sqrt{n}} \right)$
1.28	80%	$\bar{x} - E < \mu < \bar{x} + E$
1.645	90%	
1.96	95%	$n = \left(\frac{Z_c \theta}{E} \right)^2$
2.58	99%	(round up)

Checks

1) RRS

2) Indep

$-n \leq N/10$ if it is rsbl...

3) normal \rightarrow 

4) θ known

note: $PI \left(\frac{3(\bar{x} - \text{med})}{s} \right)$
or NQP

7.2 (CI w/o θ)

$$E = t_c \left(\frac{s}{\sqrt{n}} \right) \quad \bar{x} - E < \mu < \bar{x} + E$$

df = n-1 (round \downarrow if not on table)

find t_c by df + CI%

label CI's

Checks

1) RRS

2) Indep

$-n \leq N/10$ if it is rsbl...

3) normal \rightarrow 

4) θ unknown / s known

7.3 (CI w/ proportions)

$$E = Z_c \left(\sqrt{\frac{\hat{p}\hat{q}}{n}} \right) \quad (\hat{p} = r/n)$$

$$\hat{p} - E < p < \hat{p} + E$$

$n = \hat{p}\hat{q} \left(\frac{Z_c}{E} \right)^2$ if given preliminary estimate
 or $n = .25 \left(\frac{Z_c}{E} \right)^2$ if not, larger estimate

Checks

1) RRS

2) Indep

$-n \leq N/10$ if it is rsbl...

3) $np > 10,$
 $nq > 10$ thus ANAIS

On calculator

Stat \rightarrow tests \rightarrow 7.1 (Zint), 7.2 (Tint), 7.3 (1-Prop Zint)

William Gossett, 1908 Ireland, Guinness Brewery

Conclusion

If we took m samples all of size $n = x$, we expect to capture pop'n mean (μ) m on m occasions

8.1 (Test Overview)

Type I: H_0 when it's true: α

Type II: H_0 when false: β

4 ingredients of test: $H_0, H_a, CV, +$ sample stat/point estimate (+ p-value)

8.2 Part 1 (θ known)

$\alpha = 0.05$ auto-go-to when not stated

z-test	α	one-tail	two-tail
	0.01	± 2.33	± 2.58
	0.05	± 1.645	± 1.96
	0.10	± 1.28	

conclusion

- the null, the alternat, $\alpha =$
- 3x statistical evidence to suggest
- \therefore
- $p\text{-val} = \square \alpha =$

8.2 Part 2 (θ unknown)

t-test instead of z

to find t_c , use df + tail area (α)

for p-value, you will get a range

$$\bar{X} \rightarrow Z: \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$$

$$\bar{X} \rightarrow t: \frac{\bar{X} - \mu}{s/\sqrt{n}}$$

8.2 (p-values)

For Z: when 2 tail, x by 2!!

Smart has to go (compared w/ α)

"wev" or "Sev"

For t: use calculated t-value

and see where it falls on

row (use df for row, +

See what α 's it falls

between) - get a range!

Some checks as intervals

8.3 (z test for p_s)

$\hat{p} = r/n$, binomials

$$\hat{p} \rightarrow Z: \frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}}$$

Checks

1) RRS

2) indep

$n \leq N/10$ ✓ (is bin...)

3) $np > 10$,
 $na > 10$ ANAIS

6.6 Normal Appx to Binomial

• Fits binomial situation (n , and set p)

• $np > 5$ and $nq > 5$, then r has a binomial distribution that is approximated by a normal distribution with

$$\mu = np$$

$$\sigma = \sqrt{npq}$$

• more n values approaches normal distribution and can be approximated to normal

• Continuity correction

• r if left part: subtract 0.5 to obtain the corresponding x -value

$$X = r - 0.5$$

• r if right part: add 0.5 to get the corresponding x value

$$X = r + 0.5$$

example, $P(6 \leq r \leq 10)$ would be $P(5.5 \leq x \leq 10.5)$

if $P(r \leq 6)$ would be $P(x \leq 6.5)$, if $P(r < 6)$ would be $P(x < 5.5)$

★
(as histogram includes lower + upper bound)

Example: $n=40$, $p=0.5$

$np=20 > 5$, thus normal appx is justified (ANATS)

a) $nq=20 > 5$, thus normal appx is justified (ANATS)

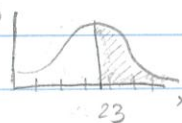
b) compute μ and σ of normal distr approximated

$$\mu = np = 20$$

$$\sigma = \sqrt{npq} = 3.16$$

c) continuity correction for $r \geq 23$ to normal variable x

23 is lower boundary, so $P(x > 22.5)$



Sampling distribution of \hat{p} :

$$\hat{p} = r/n$$

\bar{x} for binomial

$$\mu_{\hat{p}} = p$$

$$\sigma_{\hat{p}} = \sqrt{\frac{pq}{n}}$$

